

Inflation and Correlation in Claims Run-Off Triangles

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Objective

- modeling dependence **within** claims reserving triangles
- modeling dependence **between** claims reserving triangles

Joint work with

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- ▷ Michael Merz (University of Hamburg)
- ▷ Robert Salzmann (RiskLab, ETH Zurich)

Claims development triangle

accident year <i>i</i>	development year <i>j</i>																
	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16
1994	13'109	7'246	982	706	358	257	339	161	334	172	35	205	56	32	2	7	1
1995	14'457	7'581	589	487	124	74	128	50	474	12	72	63	141	286	2	10	
1996	16'075	6'597	1'081	299	154	551	29	21	16	65	98	415	280	24	27		
1997	15'682	7'782	1'001	587	477	179	44	18	65	240	7	64	4	17			
1998	16'551	7'155	921	946	473	69	168	198	220	17	6	4	7				
1999	15'439	8'357	1'070	451	822	15	21	30	559	54	18	123					
2000	14'629	7'016	1'181	773	1'393	442	42	73	55	105	14						
2001	17'585	8'703	1'335	316	396	303	77	44	766	777							
2002	17'419	8'522	1'125	695	282	434	244	157	70								
2003	16'665	8'705	1'539	702	118	132	1'969	14									
2004	15'471	8'274	1'372	1'261	593	425	84										
2005	15'103	8'290	3'416	882	370	1'122											to be predicted
2006	14'540	8'102	929	556	83												
2007	14'590	7'746	1'104	589													
2008	13'967	7'548	1'088														
2009	12'930	7'181															
2010	12'539																

- $X_{i,j}$ denotes the payments for accident year i in development year j .
- Thus, $X_{i,j}$ is paid in calendar year $k = i + j$.

Observations at time I

accident year i	development year j																
	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16
1994	13'109	7'246	982	706	358	257	339	161	334	172	35	205	56	32	2	7	1
1995	14'457	7'581	589	487	124	74	128	50	474	12	72	63	141	286	2	10	
1996	16'075	6'597	1'081	299	154	551	29	21	16	65	98	415	280	24	27		
1997	15'682	7'782	1'001	587	477	179	44	18	65	240	7	64	4	17			
1998	16'551	7'155	921	946	473	69	168	198	220	17	6	4	7				
1999	15'439	8'357	1'070	451	822	15	21	30	559	54	18	123					
2000	14'629	7'016	1'181	773	1'393	442	42	73	55	105	14						
2001	17'585	8'703	1'335	316	396	303	77	44	766	777							
2002	17'419	8'522	1'125	695	282	434	244	157	70								
2003	16'665	8'705	1'539	702	118	132	1'969	14									
2004	15'471	8'274	1'372	1'261	593	425	84										
2005	15'103	8'290	3'416	882	370	1'122											
2006	14'540	8'102	929	556	83												
2007	14'590	7'746	1'104	589													
2008	13'967	7'548	1'088														
2009	12'930	7'181															
2010	12'539																

to be predicted

- Observed payments at time $I = 2010$ (upper triangle)

$$\mathcal{D}_I = \{X_{i,j}; i + j \leq I\}.$$

Claims prediction and claims reserves

accident year <i>i</i>	development year <i>j</i>																
	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16
1994	13'109	7'246	982	706	358	257	339	161	334	172	35	205	56	32	2	7	1
1995	14'457	7'581	589	487	124	74	128	50	474	12	72	63	141	286	2	10	
1996	16'075	6'597	1'081	299	154	551	29	21	16	65	98	415	280	24	27		
1997	15'682	7'782	1'001	587	477	179	44	18	65	240	7	64	4	17			
1998	16'551	7'155	921	946	473	69	168	198	220	17	6	4	7				
1999	15'439	8'357	1'070	451	822	15	21	30	559	54	18	123					
2000	14'629	7'016	1'181	773	1'393	442	42	73	55	105	14						
2001	17'585	8'703	1'335	316	396	303	77	44	766	777							
2002	17'419	8'522	1'125	695	282	434	244	157	70								
2003	16'665	8'705	1'539	702	118	132	1'969	14									
2004	15'471	8'274	1'372	1'261	593	425	84										
2005	15'103	8'290	3'416	882	370	1'122											to be predicted
2006	14'540	8'102	929	556	83												
2007	14'590	7'746	1'104	589													
2008	13'967	7'548	1'088														
2009	12'930	7'181															
2010	12'539																

- Predict payments $\mathcal{D}_I^c = \{X_{i,j}; i + j > I\}$ (lower triangle).
- Best-estimate reserves:** $\mathcal{R}_I = \sum_{i+j>I} \mathbb{E} [X_{i,j} | \mathcal{D}_I]$.

Independence between accident years i

accident year i	development year j																
	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16
1994	13'109	7'246	982	706	358	257	339	161	334	172	35	205	56	32	2	7	1
1995	14'457	7'581	589	487	124	74	128	50	474	12	72	63	141	286	2	10	
1996	16'075	6'597	1'081	299	154	551	29	21	16	65	98	415	280	24	27		
1997	15'682	7'782	1'001	587	477	179	44	18	65	240	7	64	4	17			
1998	16'551	7'155	921	946	473	69	168	198	220	17	6	4	7				
1999	15'439	8'357	1'070	451	822	15	21	30	559	54	18	123					
2000	14'629	7'016	1'181	773	1'393	442	42	73	55	105	14						
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2002	17'419	8'522	1'125	695	282	434	244	157	70								
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2010	12'539																

- All classical stochastic claims reserving methods assume **independence** between payments $X_{i,j}$ of different accident years i .

Claims inflation and calendar year effects

accident year <i>i</i>	development year <i>j</i>																
	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16
1994	13'109	7'246	982	706	358	257	339	161	334	172	35	205	56	32	2	7	1
1995	14'457	7'581	589	487	124	74	128	50	474	12	72	63	141	286	2	10	
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2000	14'629	7'016	1'181	773	1'393	442	42	73	55	105	14						
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2008	13'967	7'548	1'088														
2009	12'930	7'181															
2010	12'539																

- Payments in different accident years i are **not independent**.
- Different portfolios $X_{i,j,n}$ for $n = 1, \dots, N \Rightarrow$ multivariate model.

Claims inflation and calendar year effects

accident year <i>i</i>	development year <i>j</i>																
	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16
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2001	17'585	8'703	1'335	316	396	303	77	44	766	777							
2002	17'419	8'522	1'125	695	282	434	244	157	70								
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2005	15'103	8'290	3'416	882	370	1'122											
2006	14'540	8'102	929	556	83												
2007	14'590	7'746	1'104	589													
2008	13'967	7'548	1'088														
2009	12'930	7'181															
2010	12'539																

	0	1	2	3	4	5	6	7	8	9	10	11	12	13
1997	2'998	5'159	4'568	4'751	4'187	2'126	1'894	332	383	553	87	179	341	95
1998	2'484	5'183	5'980	5'521	3'772	2'228	1'463	941	218	44	209	136	39	
1999	2'596	5'907	7'458	5'291	3'949	2'039	735	898	546	132	516	71		
2000	4'247	9'412	6'707	6'999	3'723	3'270	2'458	1'057	30	1'202	127			
2001	4'946	8'970	8'510	5'658	3'982	4'018	1'021	516	810	13				
2002	3'585	8'694	8'565	7'204	4'685	3'608	2'077	1'262	875					
2003	4'618	9'658	9'035	7'691	6'060	4'338	2'277	1'160						
2004	4'787	10'871	12'028	10'114	5'094	5'248	2'971							
2005	4'876	12'285	11'611	9'679	8'254	4'260								
2006	8'684	14'287	13'146	11'650	9'321									
2007	6'991	14'857	18'648	18'214										
2008	7'748	18'656	25'230											
2009	9'806	28'842												
2010	10'216													

- Payments in different accident years i are **not independent**.
- Different portfolios $X_{i,j,n}$ for $n = 1, \dots, N \Rightarrow$ **multivariate model**.

Chain-ladder model

Define cumulative payments of portfolio $n = 1, \dots, N$ by

$$C_{i,\textcolor{red}{j},n} = X_{i,0,n} + \dots + X_{i,\textcolor{red}{j},n} = \sum_{l=0}^{\textcolor{red}{j}} X_{i,\textcolor{violet}{l},n}.$$

Chain-ladder models are of multiplicative type

$$C_{i,\textcolor{red}{j},n} = C_{i,\textcolor{red}{j}-1,n} \exp \{ \xi_{i,j,n} \},$$

with individual log-link ratios

$$\xi_{i,j,n} = \log \left(\frac{C_{i,j,n}}{C_{i,j-1,n}} \right).$$

Definition of the multivariate model

For $i \in \{1, \dots, I\}$ and $j \in \{0, \dots, J\}$ we define the random vectors

$$\begin{aligned}\xi_{i,j} &= (\xi_{i,j,1}, \dots, \xi_{i,j,N})' \in \mathbb{R}^N && \text{(portfolios } \textcolor{violet}{n} \text{),} \\ \xi_i &= (\xi_{i,0}', \dots, \xi_{i,J}')' \in \mathbb{R}^a && \text{(development years } \textcolor{violet}{j} \text{),} \\ \xi &= (\xi_1', \dots, \xi_I')' \in \mathbb{R}^d && \text{(accident years } \textcolor{violet}{i} \text{),}\end{aligned}$$

with $a = N(J + 1)$ and $d = aI$.

Recall: $C_{i,j,n} = C_{i,j-1,n} \exp \{\xi_{i,j,n}\}$.

⇒ Random vector ξ contains *all* individual log-link ratios.

Bayesian multivariate chain-ladder model

Model assumptions.

- Conditionally, given $\Theta \in \mathbb{R}^a$,
 - ▷ $\xi|_{\{\Theta\}}$ is **multivariate** Gauss with covariance matrix $\Sigma \in \mathbb{R}^{d \times d}$,
 - ▷ and mean $\mathbb{E}[\xi_i | \Theta] = \Theta$ for **all** $i \in \{1, \dots, I\}$.
- The parameter has multivariate **prior** distribution $\Theta \stackrel{(d)}{\sim} \mathcal{N}(\mu, T)$.

First consequences of the model assumptions

- The mean $\mathbb{E} [\xi_i | \Theta] = \Theta$ does **not** depend on $i \in \{1, \dots, I\}$:

$$\mathbb{E} [\xi_{i,j,n} | \Theta] = \Theta_{j,n},$$

with $\Theta_{j,n}$ reflecting the **chain-ladder factor** of portfolio n for development year j .

- Covariance matrix Σ allows for **any correlation structure** between (all) individual log-link ratios in ξ , for example, for $(i, j, n) \neq (l, k, m)$

$$\text{Corr} (\xi_{i,j,n}, \xi_{l,k,m} | \Theta) = \rho > 0 \quad \text{for all } i + j = l + k.$$

These are exactly claims inflation and calendar year effects.

Bayesian multivariate chain-ladder model

Model assumptions.

- Conditionally, given $\Theta \in \mathbb{R}^a$,
 - ▷ $\xi|_{\{\Theta\}}$ is **multivariate** Gauss with covariance matrix $\Sigma \in \mathbb{R}^{d \times d}$,
 - ▷ and mean $\mathbb{E}[\xi_i | \Theta] = \Theta$ for **all** $i \in \{1, \dots, I\}$.
- The parameter has multivariate **prior** distribution $\Theta \stackrel{(d)}{\sim} \mathcal{N}(\mu, T)$.

Consequences of the model assumptions

- Appropriate choices of Σ allow to model **any correlation structure**, such as inflation, **within** the triangles and **between** the triangles:

$$\xi|_{\{\Theta\}} \stackrel{(d)}{\sim} \mathcal{N}(A\Theta, \Sigma),$$

for an appropriate matrix $A = (\mathbb{1}, \dots, \mathbb{1})' \in \mathbb{R}^{d \times a}$ (different accident years i are characterized by the same $\Theta_{j,n}$).

- The choice of the prior distribution allows to implement **parameter uncertainty** and **expert knowledge** about chain-ladder factors

$$\Theta \stackrel{(d)}{\sim} \mathcal{N}(\mu, T).$$

Predictive distribution in lower triangles

Denote by $\xi^{\mathcal{D}_I}$ the observations in the upper triangles \mathcal{D}_I and by $\xi^{\mathcal{D}_I^c}$ the unobserved components in the lower triangles \mathcal{D}_I^c .

Theorem 1 (predictive distribution).

We have posterior distribution

$$\xi^{\mathcal{D}_I^c} \mid \{\xi^{\mathcal{D}_I}\} \stackrel{(d)}{\sim} \mathcal{N}(\mu^{\text{post}}, S^{\text{post}}),$$

with explicit (credibility) formulas for μ^{post} and S^{post} .

S^{post} contains both covariance matrix Σ and parameter uncertainty T .

Claims reserves and prediction uncertainty

Corollary 2.

The ultimate claim predictor is given by

$$\widehat{C}_{i,J,n} = \mathbb{E} [C_{i,J,n} | \mathcal{D}_I] = C_{i,I-i,n} \exp \left\{ \mathbf{e}'_{i,n} \boldsymbol{\mu}^{\text{post}} + \frac{1}{2} \mathbf{e}'_{i,n} S^{\text{post}} \mathbf{e}_{i,n} \right\},$$

for appropriate projections $\mathbf{e}_{i,n}$.

- There are similar closed formula for prediction uncertainty (MSEP).
- Closed formula for the one-year claims development result (CDR) uncertainty is also available.

Summary

- We can choose
 - ★ **any correlation structure** Σ between individual log-link ratios $\xi_{i,j,n}$'s, and
 - ★ **any prior information** μ and T on the parameter space Θ ,
- and then Theorem 1 provides the **posterior distribution** of $\xi^{\mathcal{D}_I^c}$, given the observations $\xi^{\mathcal{D}_I}$ in *closed form*.
- We have closed form solutions for
 - ★ the ultimate claim predictor and the claims reserves (Corollary 2),
 - ★ MSEP for total uncertainty and one-year CDR uncertainty.
- Our model allows in **solvency models** for a

bottom-up calibration of correlation.

Case study from Merz-W.-Hashorva [1]

	0	1	2	3	4	5	6	7	8	9	10	11	12	13
1	59'966	103'186	91'360	95'012	83'741	42'513	37'882	6'649	7'669	11'061	-1'738	3'572	6'823	1'893
2	49'685	103'659	119'592	110'413	75'442	44'567	29'257	18'822	4'355	879	4'173	2'727	-776	
3	51'914	118'134	149'156	105'825	78'970	40'770	14'706	17'950	10'917	2'643	10'311	1'414		
4	84'937	188'246	134'135	139'970	74'450	65'401	49'165	21'136	596	24'048	2'548			
5	98'921	179'408	170'201	113'161	79'641	80'364	20'414	10'324	16'204	-265				
6	71'708	173'879	171'295	144'076	93'694	72'161	41'545	25'245	17'497					
7	92'350	193'157	180'707	153'816	121'196	86'753	45'547	23'202						
8	95'731	217'413	240'558	202'276	101'881	104'966	59'416							
9	97'518	245'700	232'223	193'576	165'086	85'200								
10	173'686	285'730	262'920	232'999	186'415									
11	139'821	297'137	372'968	364'270										
12	154'965	373'115	504'604											
13	196'124	576'847												
14	204'325													

	0	1	2	3	4	5	6	7	8	9	10	11	12	13
1	114'423	133'538	65'021	31'358	27'139	-377	9'889	4'477	-316	7'108	-1'035	103	209	-109
2	152'296	152'879	71'438	41'686	22'009	25'315	7'961	4'843	-113	1'593	848	4'383	-1'164	
3	144'325	162'919	106'365	50'432	55'224	7'951	8'234	1'409	2'061	669	176	977		
4	145'904	161'732	79'458	46'642	29'384	15'811	3'598	5'527	-2'484	462	-1'018			
5	170'333	171'168	92'601	36'227	11'872	18'760	3'180	3'538	948	-875				
6	189'643	171'480	85'734	61'226	18'479	13'556	7'523	1'964	88					
7	179'022	217'202	101'080	56'183	28'362	29'791	11'244	12'568						
8	205'908	210'139	104'397	45'277	34'888	30'193	17'563							
9	210'951	215'478	98'618	62'846	52'435	22'824								
10	213'426	295'796	140'211	82'259	59'209									
11	249'508	330'502	142'126	122'023										
12	258'425	427'587	229'097											
13	368'762	540'304												
14	394'997													

Case study from Merz-W.-Hashorva [1]

	0	1	2	3	4	5	6	7	8	9	10	11	12	13
1	59'966	103'186	91'360	95'012	83'741	42'513	37'882	6'649	7'669	11'061	-1'738	3'572	6'923	1'893
2	49'685	103'659	119'592	110'413	75'442	44'567	29'257	18'822	4'355	879	4'173	2'727	-776	
3	51'914	118'134	149'156	105'925	78'970	40'770	14'706	17'950	10'917	2'643	10'311	1'414		
4	84'937	188'246	134'135	139'970	74'450	65'401	49'165	21'136	596	24'048	2'548			
5	98'921	179'408	170'201	113'161	79'641	80'364	20'414	10'324	16'204	-265				
6	71'708	173'879	171'295	144'076	93'694	72'161	41'545	25'245	17'497					
7	92'350	193'157	180'707	153'816	121'196	86'753	45'547	23'202						
8	95'731	217'413	240'558	202'276	101'881	104'966	59'416							
9	97'518	245'700	232'223	193'576	165'086	85'200								
10	173'686	285'730	262'920	232'999	186'415									
11	139'821	297'137	372'968	364'270										
12	154'965	373'115	504'604											
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1	114'423	133'538	65'021	31'358	27'139	-377	9'889	4'477	-316	7'108	-1'035	103	209	-109
2	152'296	152'879	71'438	41'686	22'009	25'315	7'961	4'843	-113	1'593	848	4'383	-1'164	
3	144'325	162'919	106'365	50'432	55'224	7'951	8'234	1'409	2'061	669	176	977		
4	145'904	161'732	79'458	46'642	29'384	15'811	3'594	5'527	-2'484	462	-1'018			
5	170'333	171'168	92'601	36'227	11'872	18'760	3'180	3'538	948	-875				
6	189'643	171'480	85'734	61'226	18'479	13'556	7'523	1'964	88					
7	179'022	217'202	101'080	56'183	28'362	29'791	11'244	12'568						
8	205'908	210'139	104'397	45'277	34'888	30'193	17'563							
9	210'951	215'478	98'618	62'846	52'435	22'824								
10	213'426	295'796	140'211	82'259	59'209									
11	249'508	330'502	142'126	122'023										
12	258'425	427'587	229'097											
13	368'762	540'304												
14	394'997													

	0	1	2	3	4	5	6	7	8	9	10	11	12	13
1	114'423	133'538	65'021	31'358	27'139	-377	9'889	4'477	-316	7'108	-1'035	103	209	-109
2	152'296	152'879	71'438	41'686	22'009	25'315	7'961	4'843	-113	1'593	848	4'383	-1'164	
3	144'325	162'919	106'365	50'432	55'224	7'951	8'234	1'409	2'061	669	176	977		
4	145'904	161'732	79'458	46'642	29'384	15'811	3'598	5'527	-2'484	462	-1'018			
5	170'333	171'168	92'601	36'227	11'872	18'760	3'180	3'538	948	-875				
6	189'643	171'480	85'734	61'226	18'479	13'556	7'523	1'964	88					
7	179'022	217'202	101'080	56'183	28'362	29'791	11'244	12'568						
8	205'908	210'139	104'397	45'277	34'888	30'193	17'563							
9	210'951	215'478	98'618	62'846	52'435	22'824								
10	213'426	295'796	140'211	82'259	59'209									
11	249'508	330'502	142'126	122'023										
12	258'425	427'587	229'097											
13	368'762	540'304												
14	394'997													

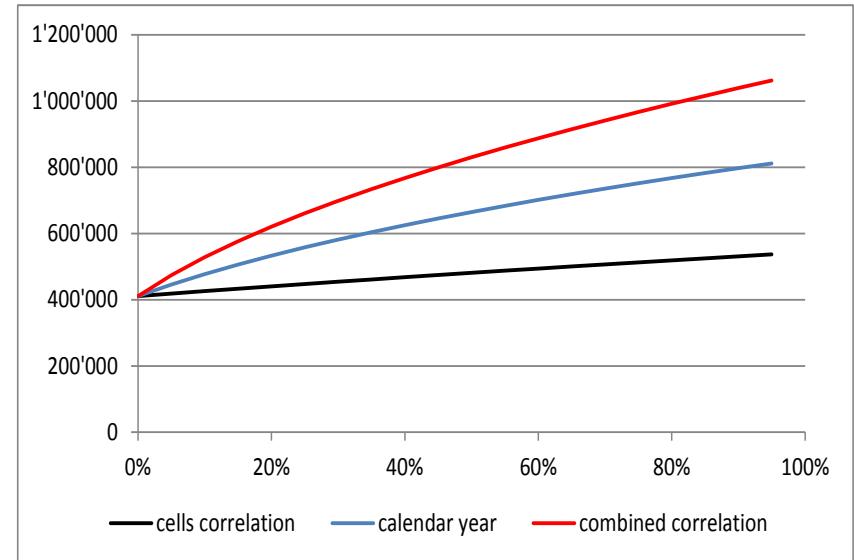
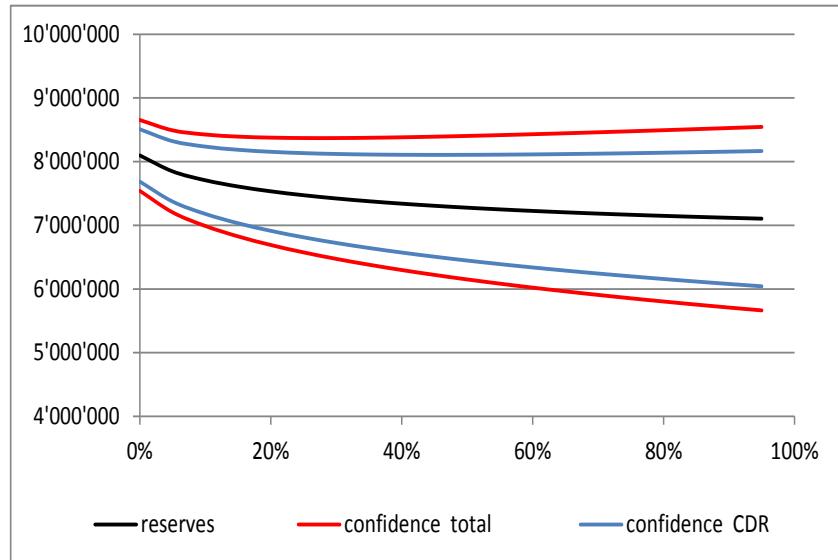
(1) point-wise cells correlation **between** triangles

(2) calendar year correlation **within** triangles

(3) combined: calendar year correlation
within and between triangles

	0	1	2	3	4	5	6	7	8	9	10	11	12	13
1	59'966	103'186	91'360	95'012	83'741	42'513	37'882	6'649	7'669	11'061	-1'738	3'572	6'823	1'893
2	49'685	102'659	119'592	110'413	75'442	44'567	29'257	18'822	4'355	879	4'173	2'727	-776	
3	51'914	118'134	149'156	105'925	78'970	40'770	14'706	17'950	10'917	2'643	10'311	1'414		
4	84'937	188'246	134'135	139'970	74'450	65'401	49'165	21'136	596	24'048	2'548			
5	98'921	179'408	170'201	113'161	79'641	80'364	20'414	10'324	16'204	-265				
6	71'708	173'879	171'295	144'076	93'694	72'161	41'545	25'245	17'497					
7	92'350	193'157	180'707	153'816	121'196	86'753	45'547	23'202						
8	95'731	217'413	240'558	202'276	101'881	104'966	59'416							
9	97'518	245'700	232'223	193'576	165'086	85'200								
10	173'686	285'730	262'920	232'999	186'415									
11	139'821	297'137	372'968	364'270										
12	154'965	373'115	504'604											
13	196'124	576'847												
14	204'325													

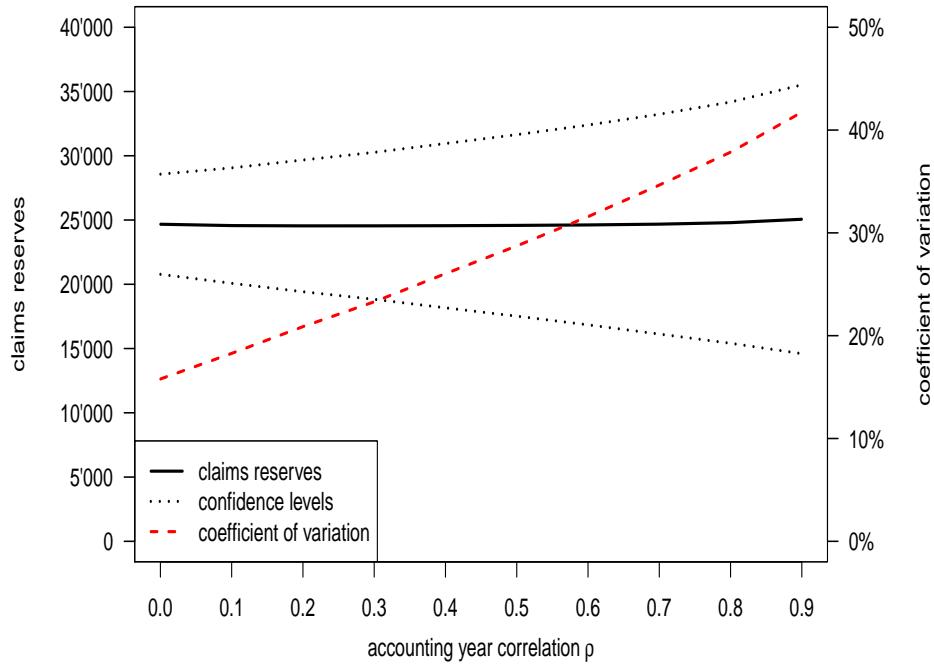
Case study from Merz-W.-Hashorva [1]



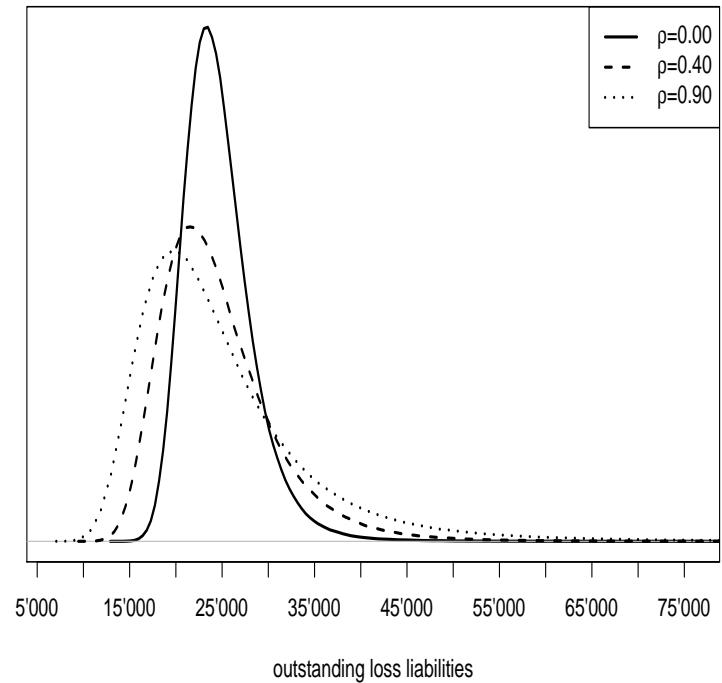
- (a) lhs: prediction and 1-std.dev. confidence interval of total uncertainty and one-year CDR uncertainty as a function of $\rho \in [0, 1]$ for combined correlation (3).
- (b) rhs: comparison (1) point-wise cells correlation, (2) calendar year correlation and (3) combined correlation for one-year CDR uncertainty and $\rho \in [0, 1]$.

Case study from Salzmann-W. [2]

claims reserves with confidence intervals



empirical density plots for different levels of ρ



(a) lhs: prediction and 1-std.dev. confidence interval of total uncertainty as a function of calendar year correlation $\rho \in [0, 1]$.

(b) rhs: posterior density of outstanding loss liabilities for different calendar year correlation choices $\rho \in \{0.0, 0.4, 0.9\}$.

Conclusions

- Dependencies and correlation have a huge effect on confidence intervals and quantiles: more case studies are needed!
- Effects of correlations are often counter-intuitive, see Conclusions 1-5 in Merz-W.-Hashorva [1].
- Our model allows in **solvency models** for a

bottom-up calibration of correlation.

- Choices of reasonable correlation structures need to be investigated: Shi-Basu-Meyers [3] and W. [4] give some advice.

References

- [1] Merz, M., Wüthrich, M.V., Hashorva, E. (2012). Dependence modelling in multivariate claims run-off triangles. To appear in *Annals Actuarial Science*.
- [2] Salzmann, R., Wüthrich, M.V. (2012). Modeling accounting year dependence in runoff triangles. To appear in *European Actuarial Journal*.
- [3] Shi, P., Basu, S., Meyers, G.G. (2012). A Bayesian log-normal model for multivariate loss reserving. *North American Actuarial Journal* 16/1, 29-51.
- [4] Wüthrich, M.V. (2012). Discussion of "A Bayesian log-normal model for multivariate loss reserving" by Shi-Basu-Meyers. To appear in *North American Actuarial Journal*.